

Problem 1. Let X be a geometric random variable with parameter $p \in (0, 1)$, i.e., with p.m.f.

$$p_X(k) = P(X = k) = (1 - p)^{k-1} p, \quad k \in \mathbb{N} = \{1, 2, 3, \dots\}.$$

- From the p.m.f. of X , find the probability that $X > n$ for $n \in \mathbb{N}$.
- Using your result from (a), show that the geometric random variable is *memoryless*, i.e., that it satisfies the property $P(X > n + m | X > n) = P(X > m)$ for any $n \in \mathbb{N}$ and $m \in \mathbb{N}$.
- Recall the meaning of the geometric random variables, and explain the reason for calling the property you proved in part (b) “memorylessness” (i.e., lack of memory).
- Can you explain the the reason for the memorylessness of geometric random variables intuitively, without any math?

Problem 2. Let the transition matrix of a Markov chain with two states be $\mathbf{P} = \begin{pmatrix} 1-a & a \\ b & 1-b \end{pmatrix}$, where a and b are numbers between 0 and 1.

- Show that $\mathbf{P}^n = \frac{1}{a+b} \begin{pmatrix} b & a \\ b & a \end{pmatrix} + \frac{(1-a-b)^n}{a+b} \begin{pmatrix} a & -a \\ -b & b \end{pmatrix}$.

Hint: One way to do this is the following. Let α_1 and α_2 be the roots of the quadratic equation $\det(\mathbf{P} - \alpha \mathbf{I}) = 0$ (where \mathbf{I} is the unit 2×2 matrix). In this problem you should obtain $\alpha_1 = 1$, $\alpha_2 = 1 - a - b$. If $\alpha_1 \neq \alpha_2$ (as in this problem), then $\mathbf{P}^n = \mathbf{C}_1 \alpha_1^n + \mathbf{C}_2 \alpha_2^n$, where \mathbf{C}_1 and \mathbf{C}_2 are 2×2 matrices whose entries can be found from

$$\mathbf{P}^0 = \mathbf{C}_1 + \mathbf{C}_2 = \mathbf{I}, \quad \mathbf{P}^1 = \mathbf{C}_1 \alpha_1 + \mathbf{C}_2 \alpha_2 = \mathbf{P}.$$

- Find $\lim_{n \rightarrow \infty} \mathbf{P}^n$.
- Use your result from part (b) to find the stationary distribution $\boldsymbol{\pi}$ of this Markov chain.

Problem 3. Let $N = \{N_t : t \geq 0\}$ be a Poisson process of intensity λ . We showed in class that the number of events occurring in the interval $(0, t]$ is

$$P(N_t = j) = \frac{(\lambda t)^j}{j!} e^{-\lambda t}, \quad j \in \mathbb{Z}_+ = \{0, 1, 2, 3, \dots\}.$$

- What type of random variable is N_t ? What are $E[N_t]$ and $\text{Var } N_t$?

Hint: Use the property of this type of random variables “for free” (i.e., you don’t need to do any computations).

- (b) If $s < t$ and $i < j$, compute the probability of the event $\{N_t = j\}$ given that the event $\{N_s = i\}$ has occurred.

Hint: In other words, you have to compute the conditional probability $P(N_t = j | N_s = i)$. This can be done without *any* calculation if you just think about the meaning of N_t ! But I do want to see your reasoning!

- (c) Again, let $s < t$ and $i < j$. Show that the probability of the event $\{N_t = j\} \cap \{N_s = i\}$ is equal to

$$P(N_t = j, N_s = i) = \frac{\lambda^j}{i!(j-i)!} (t-s)^{j-i} s^i e^{-\lambda t}.$$

Hint: Note that $P(N_t = j, N_s = i) = P(N_t = j | N_s = i) P(N_s = i)$, and use your result from part (b).

Problem 4. A *death process* is a random process that describes the number of people in a society where the only reason for changing the number of people is dying (nobody is born, there is no immigration, etc.). We say that the random process $X = \{X_t : t \geq 0\}$ is a death process with parameter μ if each person dies independently of every other person, and the probability that each person dies in one unit of time is μ (we assume that the units of time we use are much shorter than the average lifetime of the people). Clearly, the probability of a death of a person in one unit of time in a population of i people is $i\mu$ (again, we assume that the unit of time is “short”).

Here is the precise mathematical definition of a death process with parameter μ :

- the state space of the process is $\mathbb{Z}_+ = \{0, 1, 2, 3, \dots\}$;
- the process is non-increasing, i.e., if $s < t$, then $X_s \geq X_t$;
- if h is a very small positive number, then, for $j \in \mathbb{N} = \{1, 2, 3, \dots\}$,

$$P(X_{t+h} = j | X_t = i) = \begin{cases} i\mu h + o(h) & \text{if } j = i - 1, \\ 1 - i\mu h + o(h) & \text{if } j = i, \\ o(h) & \text{if } j > i \text{ or } j \leq i - 2, \end{cases}$$

and

$$P(X_{t+h} = 0 | X_t = i) = \begin{cases} \mu h + o(h) & \text{if } i = 1, \\ 1 & \text{if } i = 0, \\ o(h) & \text{if } i > 1. \end{cases}$$

Here $o(h)$ is a function satisfying $\lim_{h \rightarrow 0} \frac{o(h)}{h} = 0$;

- for $s < t$, the difference $X_t - X_s$ (equal to the number of deaths in the interval $(s, t]$) does not depend on what has happened in the time interval $(0, s]$.

In this problem you will analyze some aspects of this process.

- (a) Let $p_i(t) = P(X_t = i)$. Condition on X_t to derive the equations

$$\begin{aligned} p_0(t+h) &= \mu h p_1(t) + p_0(t) + o(h) , \\ p_j(t+h) &= (j+1)\mu h p_{j+1}(t) + (1-j\mu h)p_j(t) + o(h) , \quad j \in \mathbb{N} . \end{aligned}$$

- (b) Subtract $p_j(t)$ from the j th equation from (a), divide through by h , and take the limit $h \rightarrow 0$, to obtain the system

$$\begin{aligned} p'_0(t) &= \mu p_1(t) , \\ p'_j(t) &= (j+1)\mu p_{j+1}(t) - j\mu p_j(t) , \quad j \in \mathbb{N} . \end{aligned}$$

Let the initial condition be $X_0 = I$, where I is some positive integer.

- (c) Define the *generating function*

$$\Delta(\xi, t) := \sum_{j=0}^{\infty} p_j(t) \xi^j$$

and show that

$$\frac{\partial \Delta}{\partial \xi} = \sum_{j=0}^{\infty} j p_j(t) \xi^{j-1} = \sum_{j=1}^{\infty} j p_j(t) \xi^{j-1} , \quad \frac{\partial \Delta}{\partial t} = \sum_{j=0}^{\infty} p'_j(t) \xi^j .$$

- (d) Use the differential equations from part (b) to show that $\Delta(\xi, t)$ of the death process satisfies the partial differential equation

$$\frac{\partial \Delta}{\partial t} = \mu(1 - \xi) \frac{\partial \Delta}{\partial \xi} ,$$

and the initial condition $\Delta(\xi, 0) = \xi^I$ (where $I = X_0$ is the initial population).

Hint: Multiply the differential equation for $p'_j(t)$ by ξ^j and add all the equations.

- (e) How can the probabilities $p_j(t) = P(X_t = j)$ be expressed in terms of ξ -derivatives of $\Delta(\xi, t)$ evaluated at $\xi = 0$? Use this to find the explicit expressions for $P(X_t = 0)$ and $P(X_t = 1)$, using that the solution of the initial-value problem for the generating function posed in part (d) is

$$\Delta(\xi, t) = [1 + (\xi - 1)e^{-\mu t}]^I$$

(there is no need to derive this expression).

- (f) Show that

$$\frac{\partial \Delta}{\partial \xi}(1, t) = E[X_t] ,$$

and use this fact to find $E[X_t]$ for the death process.

- (g) Using the same idea as in part (f), express the variance of X_t in terms of derivatives of its generating function evaluated at $\xi = 1$. Use the explicit expression for $\Delta(\xi, t)$ given in (e) to find $\text{Var } X_t$ for the death process.

- (h) Radioactive decay is an example of a death process, if we think of a nucleus of the radioactive isotope as “alive” before it decays, and “dead” after that. The “half-life”, $T_{1/2}$, of a radioactive isotope is defined as the time after which only half of the initial number of nuclei of this isotope are “alive”. How is $T_{1/2}$ related to μ ?