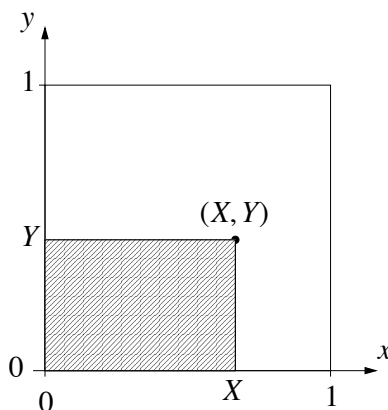


Problem 1. Let X and Y be independent random variables, each with distribution $\text{Uniform}(0, 1)$. Then the point with coordinates (X, Y) is a random vector that is uniformly distributed in the unit square, i.e., its probability density function is

$$f_{X,Y}(x, y) = \begin{cases} 1 & \text{if } 0 \leq x \leq 1, \ 0 \leq y \leq 1, \\ 0 & \text{otherwise.} \end{cases}$$

Let A be the random variable equal to the area of the rectangle with vertices at the points $(0, 0)$, $(0, Y)$, $(X, 0)$, and (X, Y) – see the figure. Clearly, A is a continuous random variable taking values in the interval $[0, 1]$.



- In the (x, y) plane, sketch the domain determined by the inequalities $0 \leq x \leq 1$, $0 \leq y \leq 1$, $xy \leq a$ (where a is a value between 0 and 1).
- Show that the cumulative distribution function of A ,

$$F_A(a) = \mathbb{P}(A \leq a) = \mathbb{P}(XY \leq a) = \iint_{xy \leq a} f_{X,Y}(x, y) \, dx \, dy,$$

is given by

$$F_A(a) = \begin{cases} 0 & \text{if } a \in (-\infty, 0] , \\ a(1 - \ln a) & \text{if } a \in (0, 1] , \\ 1 & \text{if } a \in [1, \infty) . \end{cases}$$

Hint: What you drew in part (a) will be useful.

- Find the probability density function, $f_A(a)$, of A . Be sure to specify $f_A(a)$ for all values of a .
- Determine the expected value $\mathbb{E}[A]$ of the area A of the random square with sides of length X and Y .

Problem 2. Let A and B be independent events in the sample space Ω , and let I_A and I_B be the corresponding indicator random variables:

$$I_A(\omega) = \begin{cases} 1 & \text{if } \omega \in A, \\ 0 & \text{if } \omega \notin A. \end{cases}$$

Express the following indicator random variables in terms of I_A and I_B :

$$(a) I_{A^c}; \quad (b) I_{A \cap B}; \quad (c) I_{A \cup B}.$$

In each case, explain briefly your reasoning.

Hint: The easiest way to solve this problem is to come up with some guess and then check that the guess was correct.

Problem 3. Let X and Y be independent Poisson random variables with respective parameters λ and μ , i.e.,

$$p_X(k) = e^{-\lambda} \frac{\lambda^k}{k!}, \quad p_Y(k) = e^{-\mu} \frac{\mu^k}{k!}, \quad k = 0, 1, 2, \dots$$

Show that:

- (a) $X + Y$ is Poisson with parameter $\lambda + \mu$;
- (b) the conditional distribution of X given that $X + Y = n$ is binomial, and find its parameters.

Hint: (b) Note that $\{X = k\} \cap \{X + Y = n\} = \{X = k\} \cap \{Y = n - k\}$, so that, by the independence of X and Y , $\mathbb{P}(\{X = k\} \cap \{X + Y = n\}) = \mathbb{P}(\{X = k\} \cap \{Y = n - k\}) = \mathbb{P}(X = k) \mathbb{P}(Y = n - k)$.

Problem 4. Let X be a stationary discrete-time discrete-state space Markov chain with state space S consisting of two states, 0 and 1, and let the 1-step transition probability matrix of the stochastic process be

$$\mathbf{P} = \begin{bmatrix} p_{00} & p_{01} \\ p_{10} & p_{11} \end{bmatrix} = \begin{bmatrix} \frac{1}{3} & \frac{2}{3} \\ 1 & 0 \end{bmatrix}, \quad (1)$$

where $p_{ij} = \mathbb{P}(X_{n+1} = j | X_n = i)$. Assume that you do not know the exact value of the initial value X_0 of the MC X , but you know that

$$\mathbb{P}(X_0 = 0) = \frac{1}{5}. \quad (2)$$

- (a) Find the p.m.f. p_{X_0} of the initial state X_0 of the MC.
- (b) Find $\mathbb{E}[X_0]$ and $\text{Var } X_0$.
- (c) Find the p.m.f. p_{X_1} of the state X_1 of the MC at time 1.
- (d) Find $\mathbb{E}[X_1]$.
- (e) Find the p.m.f. p_{X_2} of the state X_2 of the MC at time 2.

Problem 5. Consider the MC from Problem 4. Let

$$\rho_{ij}^{(n)} := \mathbb{P}(X_n = j, X_{n-1} \neq j, \dots, X_1 \neq j | X_0 = i) , \quad n \geq 1 , \quad i, j \in \{0, 1\}$$

be the probability of moving to state j , from the initial state i , *for the first time* at the n th transition. Recall the relations

$$p_{ij}^{(n)} = \sum_{k=1}^n \rho_{ij}^{(k)} p_{jj}^{(n-k)} , \quad (3)$$

where $p_{ij}^{(n)}$ are the entries of the n -step transition probability matrix $\mathbf{P}^{(n)}$. Assume (quite reasonably) that $\mathbf{P}^{(0)} = \mathbf{I}$ (the identity matrix).

- (a) Compute explicitly $\mathbf{P}^{(0)}$, $\mathbf{P}^{(1)}$, $\mathbf{P}^{(2)}$, and $\mathbf{P}^{(3)}$.
 - (b) Use (3) to compute the value of $\rho_{00}^{(1)}$.
 - (c) Use (3) and the value of $\rho_{00}^{(1)}$ (found in (b)) to compute the value of $\rho_{00}^{(2)}$.
 - (d) Use (3) and the values of $\rho_{00}^{(1)}$ and $\rho_{00}^{(2)}$ (found in (b) and (c)) to compute the value of $\rho_{00}^{(3)}$.
-

Food for Thought Problem 1. Prove the recursive relation (3).

Solution: Let $i \in \{0, 1\}$, $j \in \{0, 1\}$, and $n \geq 1$ be fixed. Notice that the events

$$A_k := \{ \text{the MC reaches state } j \text{ for the first time in exactly } k \text{ steps} \} .$$

Clearly, for a given n , the time k when the MC reaches state j for the first time (starting from state i) can take values $1, 2, \dots, n$. It is obvious that the events A_k form a partition of the sample space Ω (why?):

$$\bigcup_{k=1}^n A_k = \Omega , \quad A_k \cap A_{k'} = \emptyset \text{ for } k \neq k' . \quad (4)$$

Note that

$$\rho_{ij}^{(k)} = \mathbb{P}(A_k | X_0 = i) . \quad (5)$$

Using (4) and (5), we obtain

$$\begin{aligned}
p_{ij}^{(n)} &= \mathbb{P}(X_n = j \mid X_0 = i) \\
&= \mathbb{P}(\{X_n = j\} \cap \Omega \mid X_0 = i) \\
&= \mathbb{P}\left(\{X_n = j\} \cap \bigcup_{k=1}^n A_k \mid X_0 = i\right) \\
&= \mathbb{P}\left(\bigcup_{k=1}^n (\{X_n = j\} \cap A_k) \mid X_0 = i\right) \\
&= \sum_{k=1}^n \mathbb{P}(\{X_n = j\} \cap A_k \mid X_0 = i) \\
&= \sum_{k=1}^n \frac{\mathbb{P}(\{X_n = j\} \cap A_k \cap \{X_0 = i\})}{\mathbb{P}(X_0 = i)} \\
&= \sum_{k=1}^n \frac{\mathbb{P}(\{X_n = j\} \cap A_k \cap \{X_0 = i\})}{\mathbb{P}(A_k \cap \{X_0 = i\})} \frac{\mathbb{P}(A_k \cap \{X_0 = i\})}{\mathbb{P}(X_0 = i)} \\
&= \sum_{k=1}^n \mathbb{P}(X_n = j \mid A_k \cap \{X_0 = i\}) \mathbb{P}(A_k \mid X_0 = i) \\
&= \sum_{k=1}^n \mathbb{P}(X_n = j \mid A_k) \mathbb{P}(A_k \mid X_0 = i) \\
&= \sum_{k=1}^n \mathbb{P}(X_n = j \mid X_k = j) \mathbb{P}(A_k \mid X_0 = i) \\
&= \sum_{k=1}^n \mathbb{P}(X_{n-k} = j \mid X_0 = j) \mathbb{P}(A_k \mid X_0 = i) \\
&= \sum_{k=1}^n p_{jj}^{(n-k)} \rho_{ij}^{(k)} .
\end{aligned}$$